

M. Math Exam

Question 1

For a fixed j , we need to find $P_{ij} = P(X_{n+1} = i | X_n = j)$ for all i . We have $P(X_n = j + 1 | X_{n-1} = j) = p$, $P(X_n = 0 | X_{n-1} = j) = p$ and $P(X_n = k | X_{n-1} = j) = 0$ for all other values of k .

If T_0 denotes the time of return to zero starting from 0, we have that $T_0 = k$ occurs when we have all successes in the first $k - 1$ tosses and the k^{th} toss is a failure. Therefore $P_0(T_0 = k) = p^{k-1}(1-p)$, for $k \geq 1$, which is a Geometric distribution with parameter p . This implies that $P_1(T_1 < \infty) = 1$. So by Proposition 6.3.5 of Athreya and Sunder 2008, we have that 0 is recurrent.

Also, from 0 it is possible to reach any state $k \geq 1$ with positive probability and therefore the Markov chain is irreducible. Thus all states are recurrent.

Question 2

(i) If $y = x$, then we have by Proposition 6.3.14 of Athreya and Sunder 2008 that x is transient and $\sum_n P_{xx}^{(n)} < \infty$. Here $P_{xx}^{(n)} = P_0(X_n = x)$. But we also have

$$\sum_n P_{xx}^{(n)} = \sum_n E_x(\mathbf{1}(X_n = x)) = E_x \sum_n \mathbf{1}(X_n = x) \quad (1)$$

and the final term is the expected number of visits to x starting from x . In the above $\mathbf{1}(\cdot)$ refers to the indicator function. Using Fubini's theorem gives the necessary interchange.

If $y \neq x$, and suppose there is a positive probability of eventually reaching y from x . Starting from y , we then again have expected finite number of visits to y . Thus again $G(x, y) < \infty$ in this case.

Finally, if $y \neq x$, and suppose there is zero probability of reaching y from x , we have $G(x, y) = 0$ in this case.

(ii) As in (1), we have

$$G(x, y) = \sum_n P_{xy}^{(n)}.$$

Since $G(x, y) < \infty$, we have $\lim_n P_{xy}^{(n)} = 0$.

(iii) Let x_1, \dots, x_t be the states of the Markov chain and suppose they are all transient. We then have

$$\lim_n P_{x_1, x_i}^{(n)} = 0 \text{ for all } i. \quad (2)$$

But $\sum_{i=1}^t P_{x_1, x_i}^{(n)} = 1$ since starting from x_1 , the Markov chain must be in one of the t states at time n . Allowing $n \rightarrow \infty$, we have $\sum_{i=1}^t \lim_n P_{x_1, x_i}^{(n)} = 1$ and here we interchange limits since the sum is finite. But thus contradicts (2).

Question 3

(i) We have $P_{ij} = P(X_{n+1} = j | X_n = i)$ and $P_{ij}^2 = \sum_k P_{ik} P_{kj}$. Also,

$$\begin{aligned} & P(X_{2n+2} = j | X_{2n} = i) \\ &= \sum_k P(X_{2n+2} = j | X_{2n+1} = k, X_{2n} = i) P(X_{2n+1} = k | X_{2n} = i) \\ &= \sum_k P(X_{2n+2} = j | X_{2n+1} = k) P(X_{2n+1} = k | X_{2n} = i) \\ &= \sum_k P_{ik} P_{kj} \end{aligned} \quad (3)$$

where (3) is due to the Markovian nature. Thus P_{ij}^2 is the transition probability matrix for $\{X_{2n}\}_n$.

(ii) Let t be the number of states. Since π is a stationary distribution, we have

$$\sum_{i=1}^t \pi(i) P_{ij} = \pi(j) \text{ for all } i, j. \quad (4)$$

Using this we have

$$\begin{aligned}
\sum_{i=1}^t \pi(i) P_{ij}^2 &= \sum_{i=1}^t \pi(i) \sum_{k=1}^t P_{ik} P_{kj} \\
&= \sum_{k=1}^t \sum_{i=1}^t \pi(i) P_{ik} P_{kj} \\
&= \sum_{k=1}^t \pi(k) P_{kj} = \pi(j)
\end{aligned}$$

Thus π is also a stationary distribution for P^2 .

(iii) No. Take state space to be $\{0, 1\}$, $P(X_{n+1} = 1|X_n = 0) = 1$ and $P(X_{n+1} = 0|X_n = 1) = 1$. Starting from 0, the Markov chain oscillates between 0 and 1 at alternate times. Here $\pi = (1, 0)$ is a stationary distribution for P^2 but not for P .

Question 4

(i) We condition on the first step X_1 , to get

$$P_x(T_y = n+1) = \sum_{z \neq y} P_x(T_y = n+1|X_1 = z) P_x(X_1 = z) = \sum_{z \neq y} P_z(T_y = n) P_{xz} \quad (5)$$

For the second equality we argue this way: if in the first step we reach $z \neq y$, then n time units are needed to reach y from z . From Markov property the equation then follows.

(ii) We have

$$\rho_{xy} = P_x(T_y < \infty) = P_x(T_y = 1) + \sum_{k \geq 2} P_x(T_y = k) = P_{xy} + \sum_{k \geq 2} P_x(T_y = k).$$

For the second term we proceed as follows. Summing (5) over $n \geq 1$, we have

$$\sum_{n \geq 1} P_x(T_y = n+1) = \sum_{z \neq y} \sum_{n \geq 1} P_z(T_y = n) P_{xz} = \sum_{z \neq y} P_{xz} \rho_{zy}$$

since the inner summation in the middle term is precisely the probability of eventually reaching y from z . In the above, the iterated sums are not necessarily finite and therefore not directly interchangeable. As before, using Fubini's theorem gives the necessary interchange.